<u>UMC</u>: Unfinished Business

<u>Donald L. Smith (</u>ANL, retired) Roberto Capote (IAEA) Denise Neudecker (LANL)



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UMC Unified Monte Carlo

The Concept

- Assume that knowledge about a set of nuclear observable parameters employed in nuclear system analyses can be represented by a multi-variate master probability function.
- This function should be **constructed** by incorporating the best available information from **theory** <u>and</u> **experiments**.
- The master function is then sampled using Monte Carlo methods to generate a Markov Chain of random observable parameter vectors that ultimately can be employed for a variety of practical applications such as generating evaluations and analyzing the behavior of derived nuclear system parameters.

Master Probability Density Function

Bayes Theorem: $p(\mathbf{x} | \mathcal{T}, \mathcal{E}) = p_0(\mathbf{x} | \mathcal{T}) L(\mathbf{x} | \mathcal{T}, \mathcal{E})$

- "3" signifies prior information based on theory (modeling).
- "&" denotes <u>independent</u> information from <u>experiments</u> that serves to <u>improve</u> (or augment) prior theoretical knowledge *S*.
- "x" represents random vectors corresponding to possible values of the nuclear observables (e.g., cross sections).
- The prior probability function p₀ is based on theory, while likelihood L is a probability function that quantifies the consistency of data from theory and experiments used to construct the master (posterior) probability function p(x | S, E).

Prior Probability Function $p_o(\mathbf{x} | \mathcal{S})$

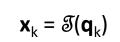
- S is a complicated algorithm that maps theoretical model parameters q to calculated observables x, i.e., x = S(q).
- By applying Monte Carlo techniques, a Markov Chain of vectors **q**_k can be generated by random sampling of parameters in space S(**q**) governed by a probability function *r*₀(**q**). Usually, mean values **q**₀ and covariance matrix V_q are specified. Then, <u>Maximum Entropy</u> suggests *r*₀ should be a normal probability function.
- A Markov chain of values \mathbf{x}_k in observables space $S(\mathbf{x})$ is generated by Monte Carlo sampling according to $\mathbf{x}_k = \mathcal{J}(\mathbf{q}_k)$.
- The collection {x_k} reflects the prior probability function p₀, but rarely (if ever) can p₀ be expressed explicitly as an analytical function that can be sampled in a conventional way!

Topology Issues

The schematic diagram shows mapping from space $S{\mathbf{q}}$ to space $S{\mathbf{x}}$ by the theoretical (model) algorithm \mathcal{S} . The shaded areas denote regions of non-negligible probability for r_0 (green) and p_0 (blue).

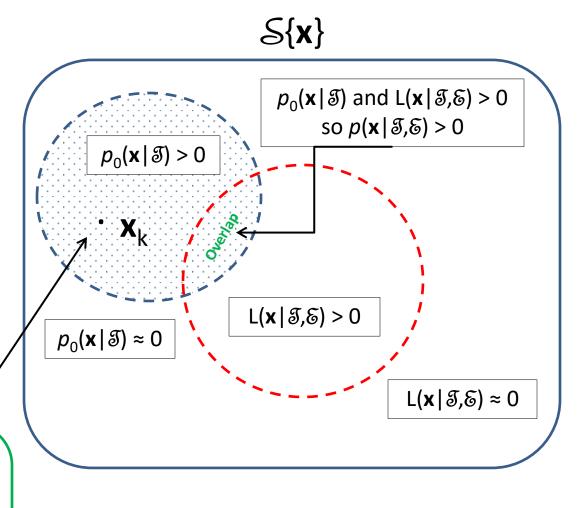
 $r_0({\bf q}) > 0$

 \mathbf{q}_k



 $r_0(\mathbf{q})\approx 0$

S{**q**}



The region enclosed by a **red** dashed circle indicates that portion of space $\mathcal{S}\{\mathbf{x}\}$ where the **likelihood** function $L(\mathbf{x} | \mathcal{S}, \mathcal{S})$ is **non-negligible**. In the region labeled **Overlap**, where "blue" and "red" dashed circles intersect, the **master** (posterior) **function** is also **non-negligible**.

<u>UMC-G</u>: Analytical Approximation to $p_0(\mathbf{x} | \mathcal{S})$

D.L. Smith, *Proceedings of AccApp'07*, Pocatello, ID, July 29 – August 2, 2007, Amer. Nucl. Soc. , p. 736.

- The collection of K calculated observable parameter vectors $\{\mathbf{x}_k\}$ generated by Monte Carlo (see preceding two slides), according to the mapping $\mathbf{x}_k = \mathcal{S}(\mathbf{q}_k)$, is used to calculate mean values \mathbf{x}_0 and covariance matrix \mathbf{V}_x via the formulas: $\mathbf{x}_{0i} \approx (\sum_{k=1,K} \mathbf{x}_{ik}) / K$ and $(\mathbf{V}_x)_{ij} \approx [(\sum_{k=1,K} \mathbf{x}_{ik} \mathbf{x}_{jk}) / K] - \mathbf{x}_{0i} \mathbf{x}_{0j}$ (K is very large).
- The "true" prior probability function p₀(x | 𝔅) typically is approximated by a multi-variate normal probability function given by:
 p₀(x | 𝔅) ≈ C exp {-(½)[(x x₀)^T V_x⁻¹ (x x₀)]} (C is a normalization constant).

<u>Advantage</u>: A lengthy Markov Chain of sample values is thus replaced by an analytical approximation having the same mean values and covariance matrix. This yields a master (posterior) function $p(\mathbf{x} | \mathcal{T}, \mathcal{E})$ that can be sampled readily by conventional Monte Carlo methods, e.g., "Brute Force" or "Metropolis-Hastings".

<u>Disadvantage</u>: This approximation discards all information pertaining to higherorder distribution moments inherent in the Monte Carlo generated Markov chain $\{\mathbf{x}_k\}$. This rejection of information can lead to significant biases in cases where non-linear effects and distribution skewness and kurtosis are present.

<u>UMC-B</u>: Information in $p_o(\mathbf{x} | \mathcal{S})$ is Preserved

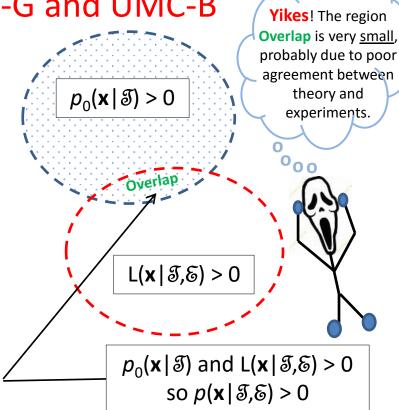
R. Capote et al., *Proceedings of ISRD-14*, Breton Woods, NH, May 22 – 27, 2011, ASTM STP-1550, p. 179.

- The collection of K calculated observable parameter vectors {x_k} generated by Monte Carlo (shown in two earlier slides), according to the mapping x_k = S(q_k), is preserved. Thus, no information on higher-order moments of p₀ is discarded.
- For each \mathbf{x}_k , a scalar weighting factor $\boldsymbol{\omega}_k$ is generated according to the expression: $\boldsymbol{\omega}_k = L(\mathbf{x}_k | \mathcal{J}, \mathcal{E})$. Thus, the worth that is assigned to each MC sampled parameter vector \mathbf{x}_k is based on its consistency with available experimental data, as reflected in the likelihood function.
- For <u>very</u> large K, it is assumed that mean values and covariance matrix for the master (posterior) probability function $p(\mathbf{x} | \mathcal{S}, \mathcal{E})$ are estimated from: $x_{0i} \approx (\sum_{k=1,K} \omega_k x_{ik}) / (\sum_{k=1,K} \omega_k)$ and $(\mathbf{V}_x)_{ij} \approx [(\sum_{k=1,K} \omega_k x_{ik} x_{jk}) / (\sum_{k=1,K} \omega_k)] - x_{0i} x_{0j}$
- The Markov Chain for UMC-B thus consists of the set of pairs $\{x_k, \omega_k\}$. These values can be used for nuclear systems applications as well as evaluations.

<u>Advantage</u>: All information in function $p_0(x \mid S)$ is clearly preserved, including that related to non-linearity as well as the distribution skewness and kurtosis.

A Closer Look at UMC-G and UMC-B

- The areas enclosed by "blue" and "red" dashed circles, respectively, indicate regions of nonnegligible probability for the model-generated prior probability and the likelihood function that quantifies the consistency of theory and experiment.
- The small region **Overlap** of these two circles is indicative of data inconsistency. Such an outcome could have potentially negative implications for an application of the UMC-B method (e.g., **limited** or **biased sampling** of the sparsely sampled region **Overlap**). Statistical **inadequacy** is **not a problem** in applying the **UMC-G** approach, but it can suffer from significant bias effects due to the explicit rejection of higher-order moments of $p_0(\mathbf{x} | \mathcal{S})$. Data inconsistency between theory and experiment will inevitably lead to evaluations and system analysis results that are very questionable and thus inherently unreliable.



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- Further investigation of the mentioned sampling issues for the region Overlap in the UMC-B approach is warranted.
- Detailed inter-comparisons of GLS, UMC-G, and UMC-B predictions for extreme cases and inconsistent data are needed.



- The original UMC-G formulation discards potentially valuable information about higher moments of the prior probability p₀(x | S) that is reflected in the Markov Chain of vectors {x_k} generated by Monte Carlo sampling.
- It is **unlikely** that the **moments** of p_0 of **higher order** than mean values, covariances, skewness, and kurtosis will **affect applications** significantly.
- The moments of p₀ (mean values, covariances, skewness, and kurtosis) can be estimated using the collection of sample vectors {x_k}.

Suggestion: Perhaps analytical functions might be found whose parameters can be adjusted to approximate the distribution moments deduced from $\{\mathbf{x}_k\}$. It could then be employed to serve as a surrogate for the master (posterior) probability function $p(\mathbf{x} | \mathcal{S}, \mathcal{S})$, and it would then be sampled using conventional Monte Carlo methods.

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- Investigate the structure of realistic MCgenerated distributions p₀(**x** | S) with the intent of quantifying typical mean values, covariances, skewness, and kurtosis.
- Identify families of analytical mathematical functions that might serve as surrogates for representing the MC-generated distributions p₀(**x** | S) with greater fidelity than using a simple normal distribution (as in UMC-G).

Thoughts on Likelihood Functions L(x | *ଶ*, &)

- Available experimental data are usually comprised of mean values and (far less often) covariances. Therefore, comparisons between theoretically calculated observables and experimental observables should involve at most mean values and covariances.
- Consequently, the likelihood function L(x | S, E), in accordance with Maximum Entropy, should be an appropriately constructed normal probability function. In particular, it should have the form:

 $L(\mathbf{x} | \mathcal{J}, \mathcal{E}) = C \exp \{-(\frac{1}{2})[(\mathbf{y} - \mathbf{y}_{E})^{T} \mathbf{V}_{E}^{-1} (\mathbf{y} - \mathbf{y}_{E})]\}$ (C is a normalization constant)

<u>Note</u>: \mathbf{y}_{E} is an **experimental data vector** with **covariance matrix** \mathbf{V}_{E} . Furthermore, $\mathbf{y} = \mathbf{f}(\mathbf{x})$, since what is measured (\mathbf{y}) may not correspond directly to the **observable parameters** (\mathbf{x}) that are being considered. The function collection " \mathbf{f} " establishes how \mathbf{x} and \mathbf{y} are related.

It may be very difficult to construct a rigorous likelihood function L(x | 𝔅,𝔅) in any given situation due to one or more of the following limitations: i) incomplete data, ii) discrepant (wrong) data, iii) weak sensitivity relationships between the data (y) and parameters of interest (x), and iv) excessive computational overhead.

<u>Alternatives</u>: Because of these limitations, some investigators (notably A. Koning and D. Rochman) for pragmatic reasons have investigated using simpler alternative likelihood functions $L(\mathbf{x} | \mathcal{T}, \mathcal{E})$.

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- The impact of experimental data quality and availability on applications of UMC needs to be investigated thoroughly.
- Improve experimental covariance data.

The UMC Approach at a Crossroads?

→ There are unresolved technical issues and unanswered questions.
The way forward to further develop UMC must be clarified.

- Would UMC-G and UMC-B be truly comparable if p₀(x | 3) could be expressed exactly as an analytical function?
- Can more sophisticated analytical function approximations to a MC prior than the normal distribution be found (e.g., UMC-G Plus)?
- Are the available **experimental data** sufficiently **accurate** and **comprehensive** to be useful in practice for **applying UMC**?
- Can better theoretical models be developed to reduce the discrepancies between theory and quality experimental data?
- If not, can model-defects formalisms be developed as practical measures to cope with model vs. experimental data discrepancies?
- How much extra "value" does UMC contribute, compared with GLSQ, to justify the additional effort and computational burden?



