

$n+^{239}\text{Pu}$ covariance evaluation at LLNL

mini-CSEWG, Livermore Valley Open Campus, April 2023

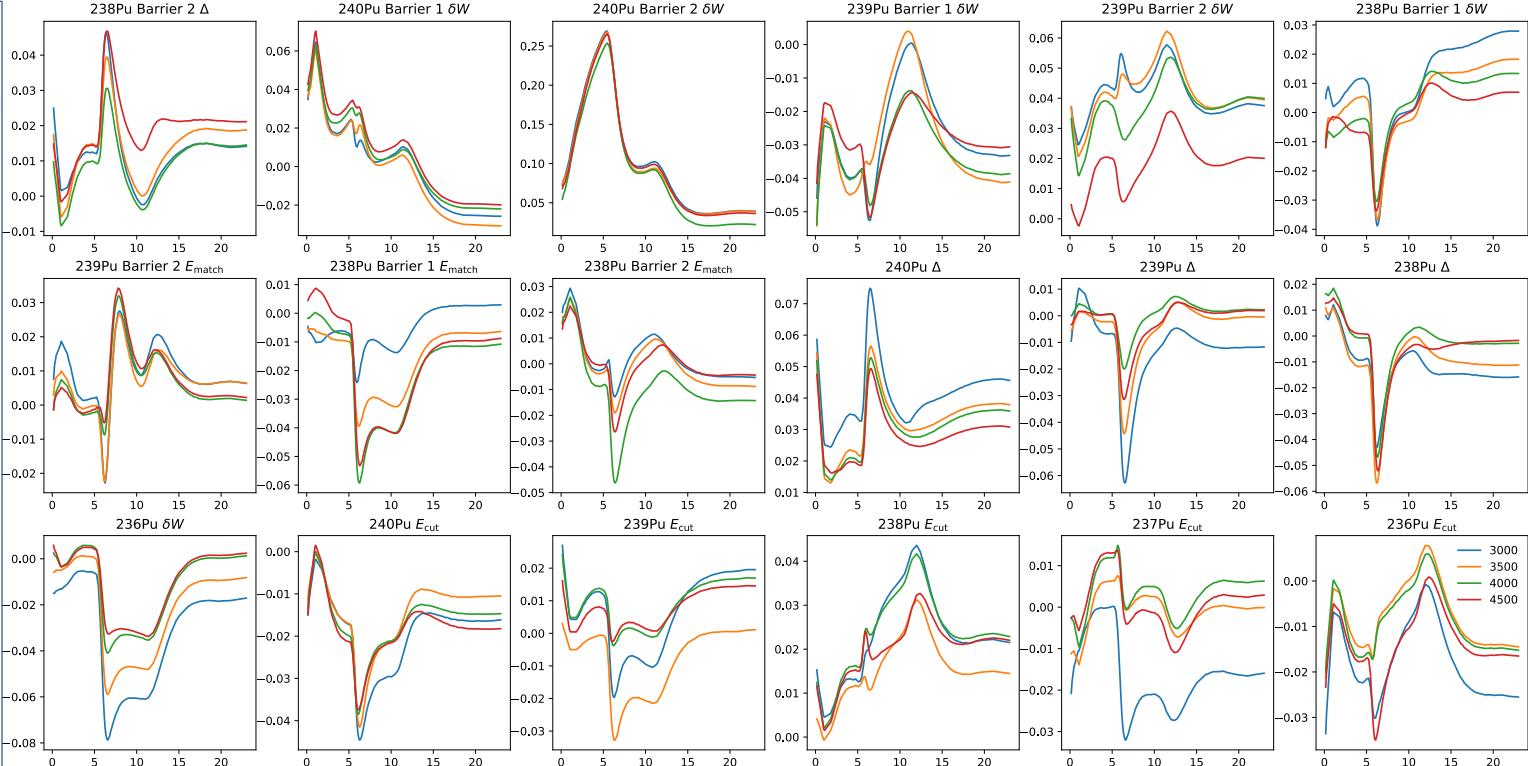
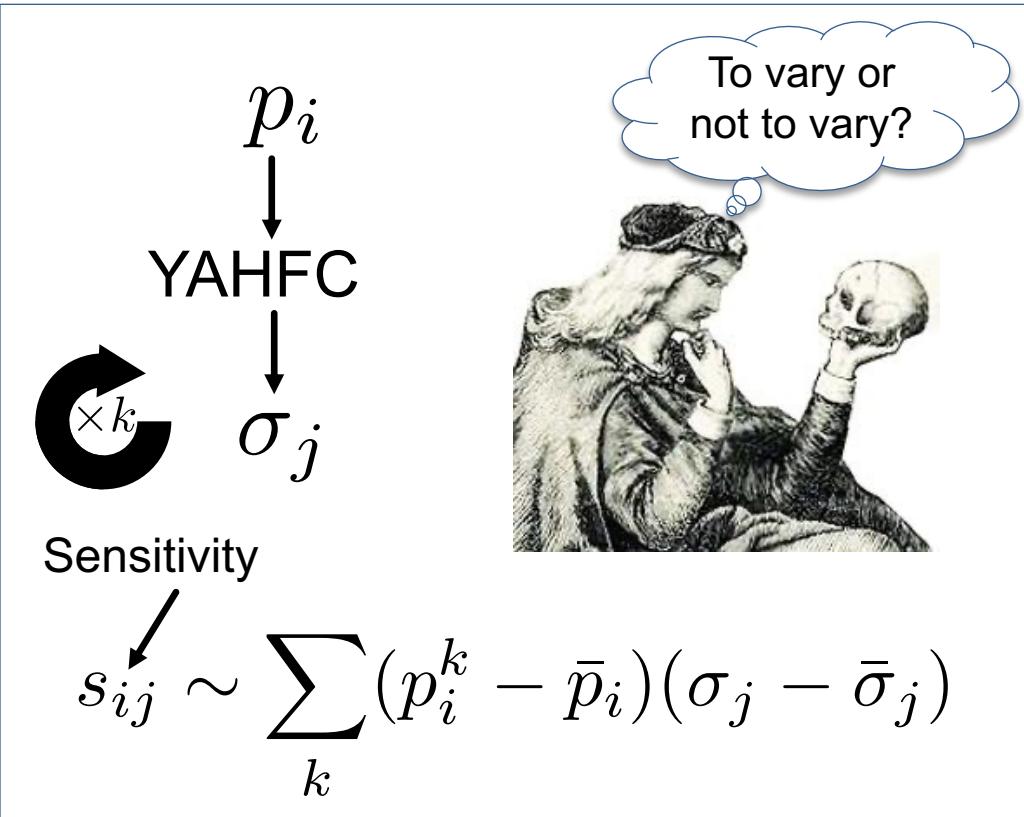
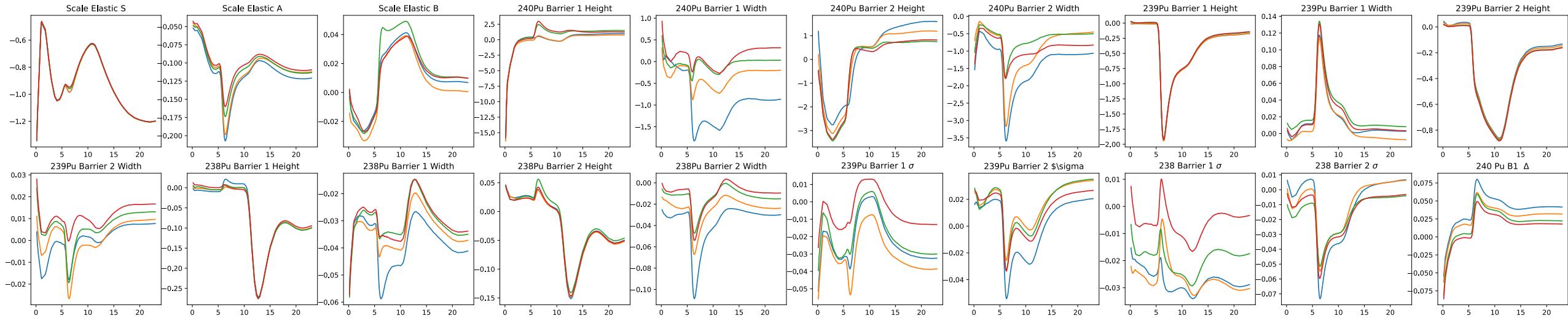
K. Kravvaris & K. Wendt



Extracting covariances from model-driven evaluations.

- Calculations using LLNL-developed YAHFC (yet another Hauser-Feshbach code.)
- Multiple parameters can be varied and are thus fitted to experimental data.
- Ultimate goal is the Markov-chain Monte-Carlo evaluation of the model covariances.
- YAHFC runs take significant amount of time, developed parallel version (not discussed here.)
- Explore options for quick re-evaluation of covariances in the wake of new experimental data.

 [LLNL / Yet-Another-Hauser-Feshbach-Code](https://github.com/LLNL/Yet-Another-Hauser-Feshbach-Code) Public
<https://github.com/LLNL/Yet-Another-Hauser-Feshbach-Code>



Kalman filter (following NDS 109 (2008) 2752–2761)

Sensitivity Matrix

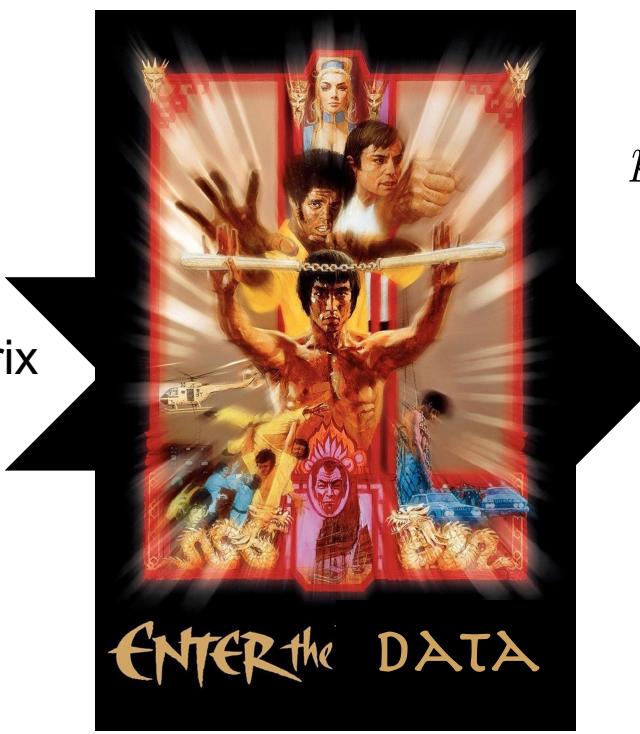
$$S_{ij} = \frac{\partial \sigma_i}{\partial p_j}$$

Parameter (p) covariance matrix
(Assumed diagonal)

Model covariance

$$C_0 = S P_0 S^T$$

$$\sigma_i \equiv \sigma(E_i, \mathbf{p})$$



Updated YAHFC parameter values

$$p_{n+1} = p_n + P_n S^T Q_{n+1} (\sigma_{n+1}^{\text{exp}} - \sigma(\mathbf{p}_n))$$

$$P_{n+1} = P_n - P_n S^T Q_{n+1} S P_n$$

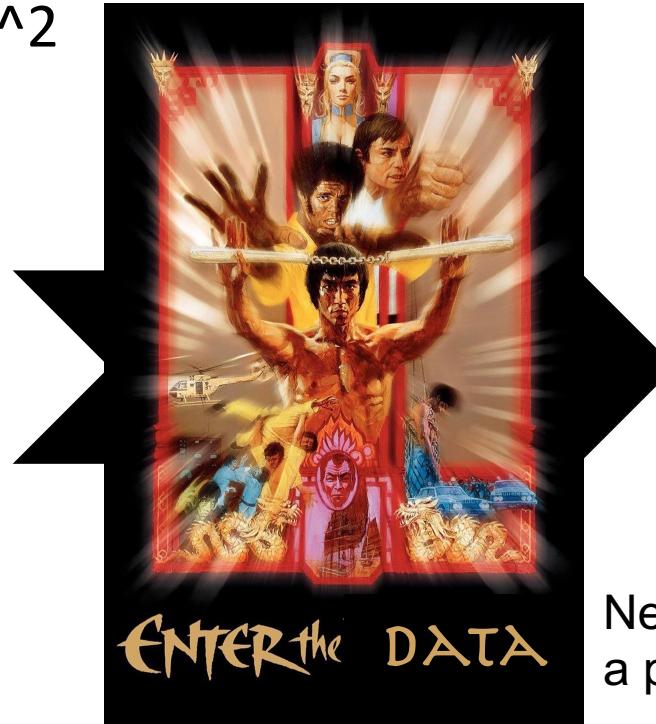
$$Q_{n+1} = (C_n + C_{n+1}^{\text{exp}})$$

$$C_{n+1} = S P_{n+1} S^T$$

Updated covariances

Backwards-Forwards Monte Carlo.

- Start from some “guess” parametrization of YAHFC (eye^2 fit)
- Draw a bunch of samples that follow a “guess” covariance matrix



- Compute χ^2 for each parametrization
- New “mean” parametrization is the Bayesian model average of the parameters with weights

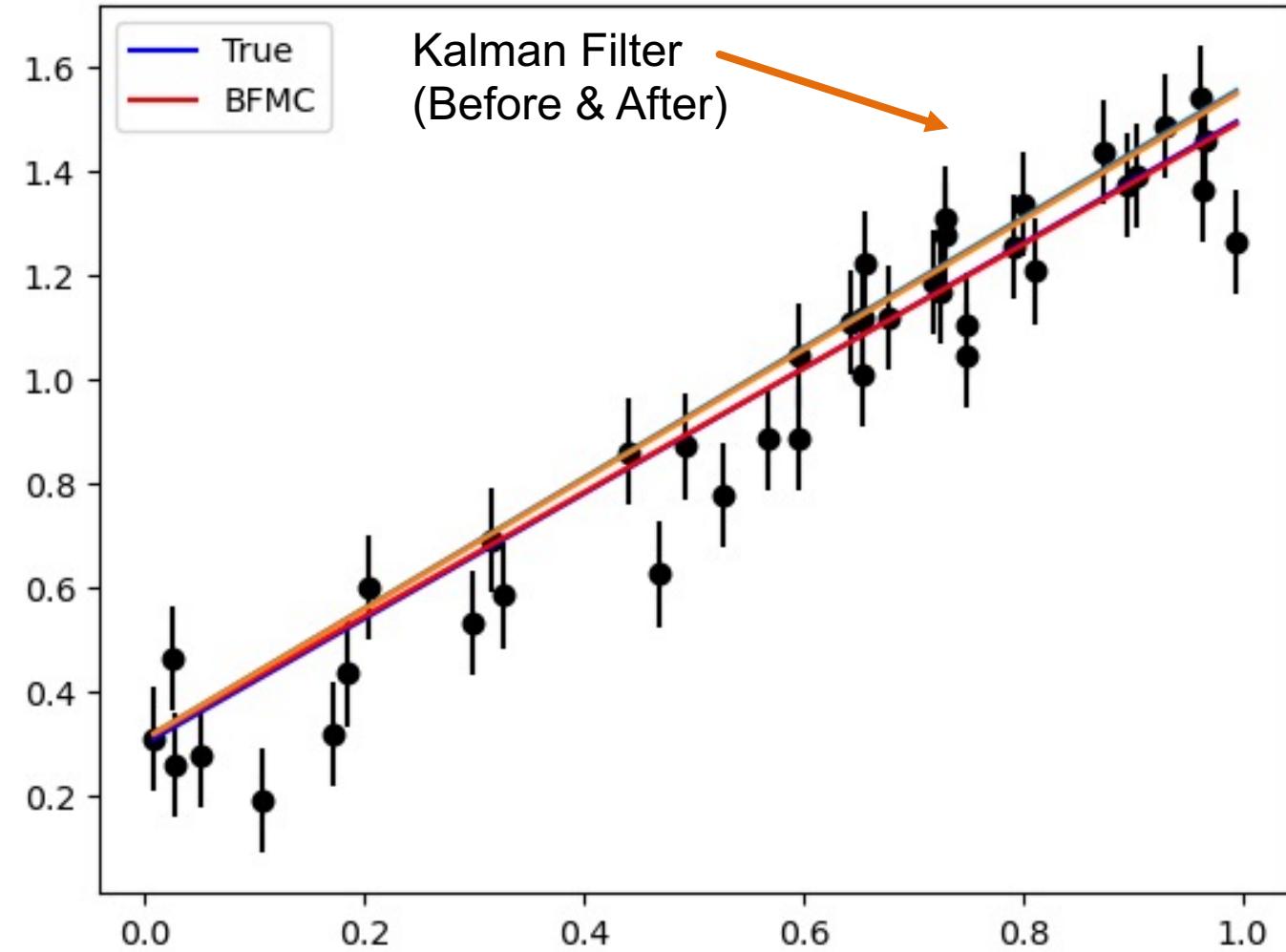
$$w_i = C \exp \left[- \left(\frac{\chi^2}{\chi_{\min}^2} \right)^2 \right]$$

New probability for a parametrization:

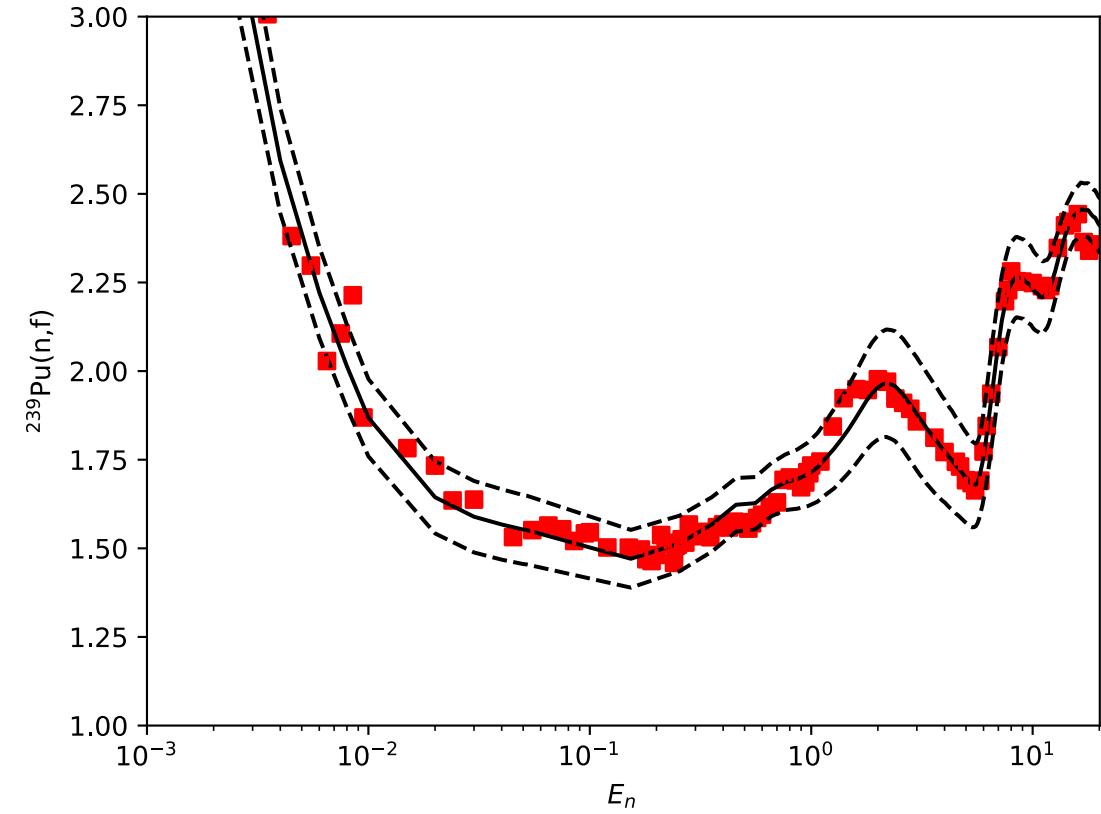
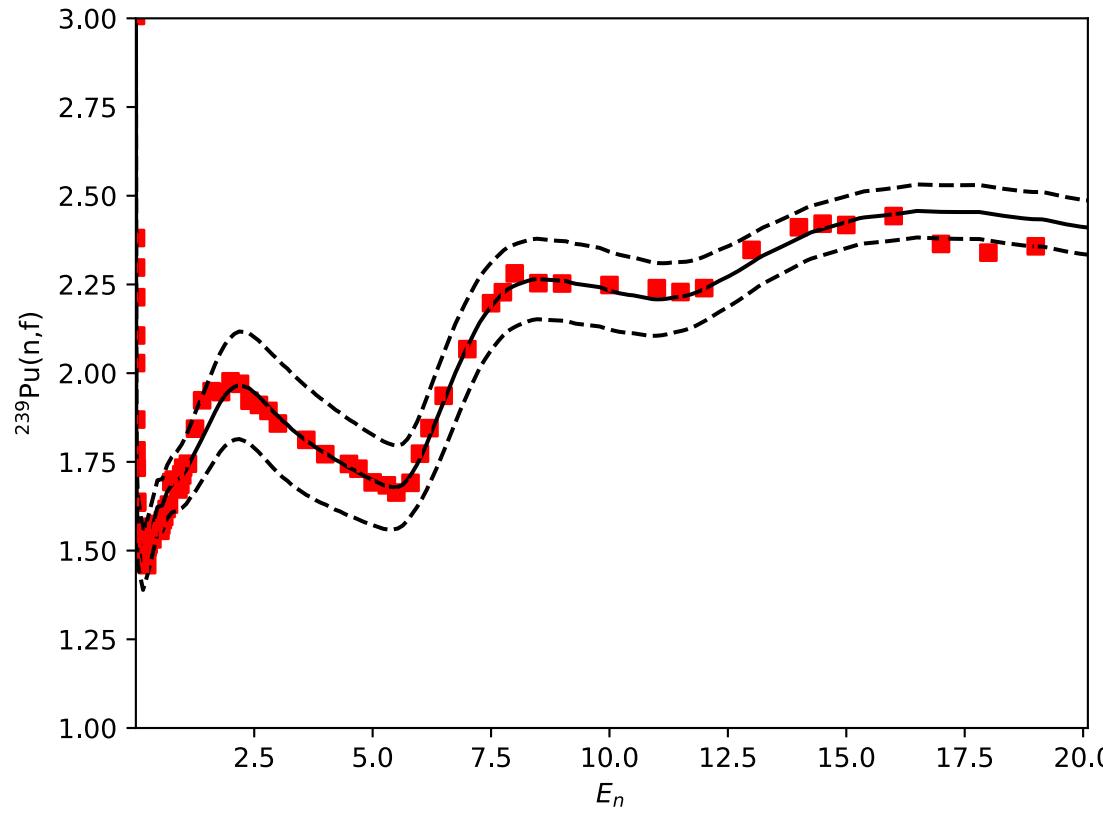
$$P(\mathbf{p}) = \sqrt{P_1} G + p_1$$

Which way to choose? BFMC seems to be performing better in toy problem.

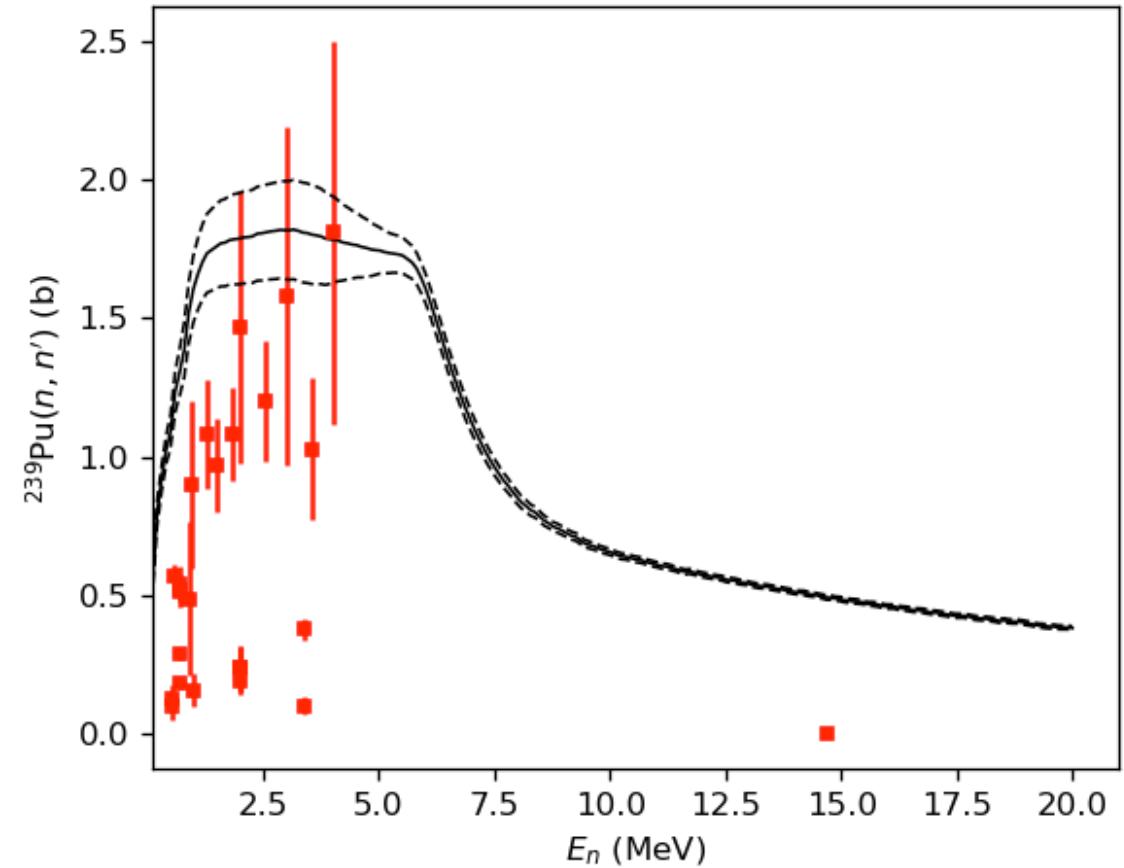
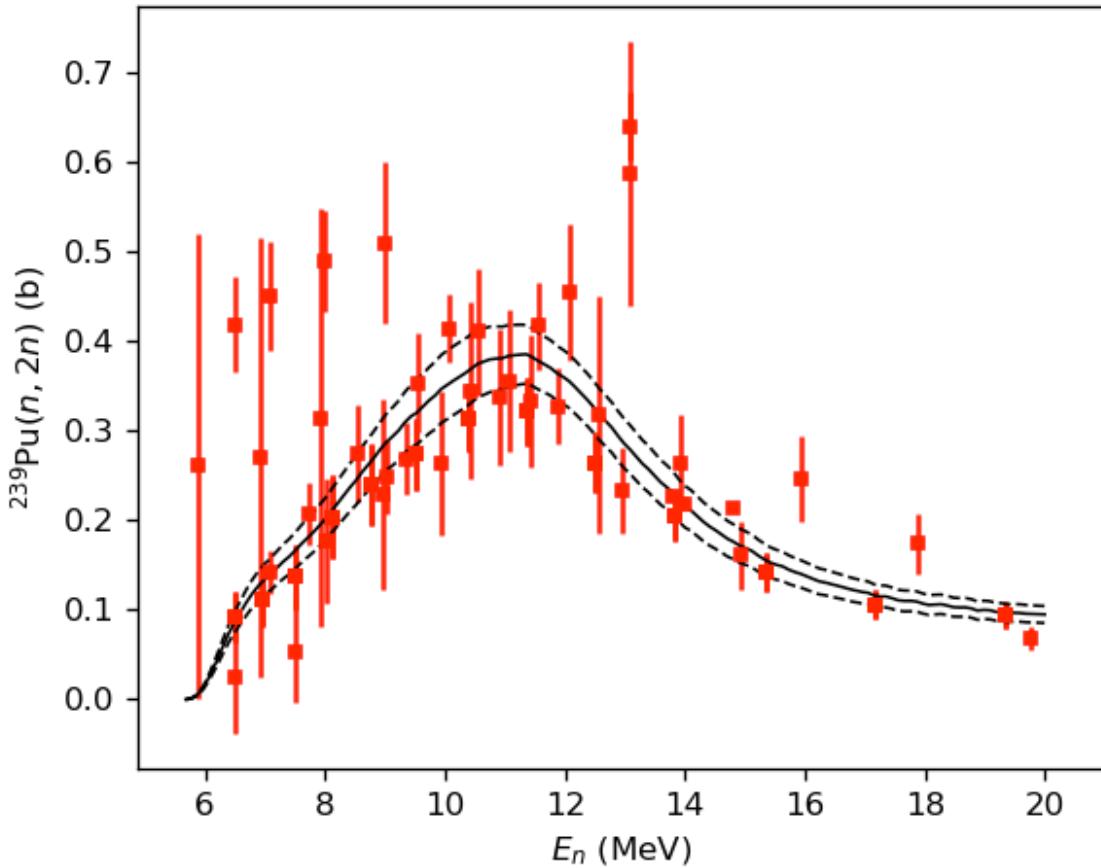
- Fit line to data with random normal error.
- Both ways are responsive when it comes to introducing new data.
- Unscented Kalman filter could be next?



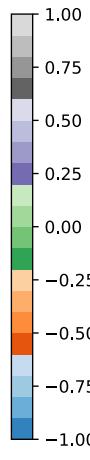
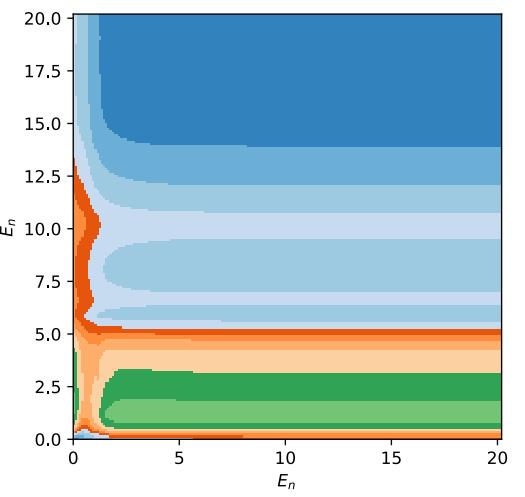
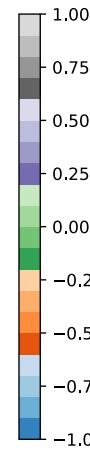
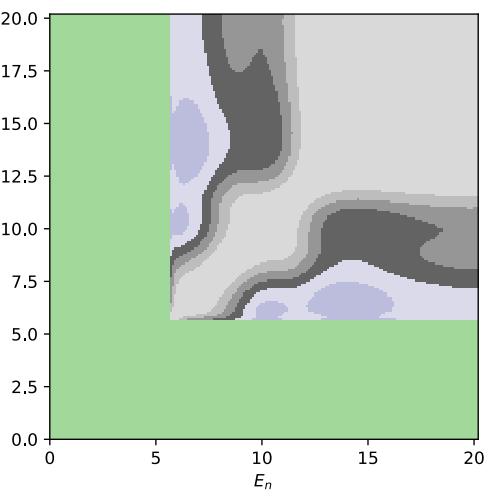
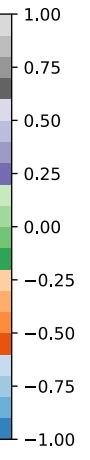
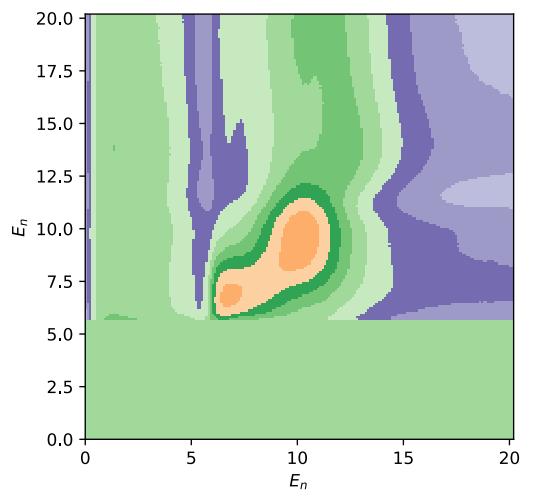
BFMC applied to the YAHFC n+P9 evaluation.



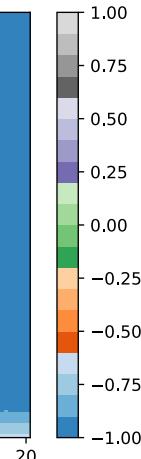
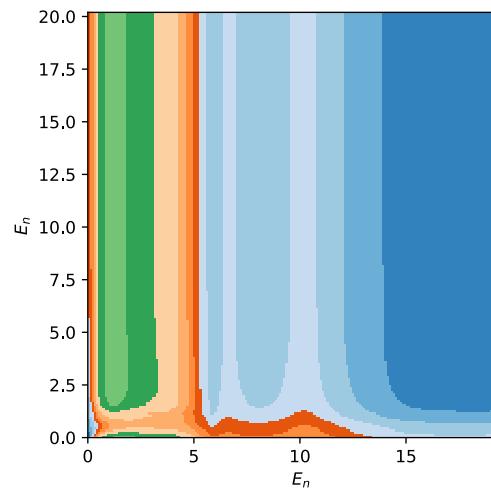
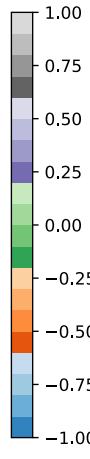
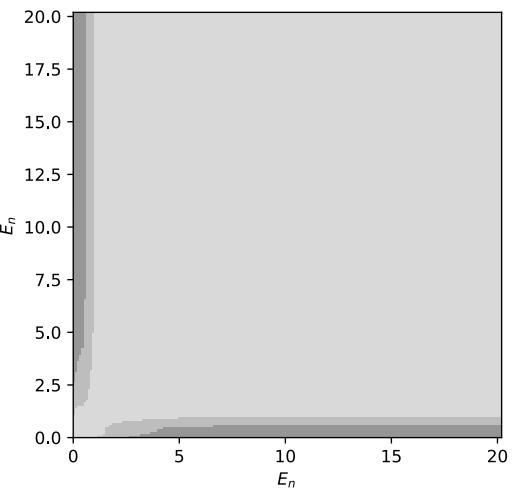
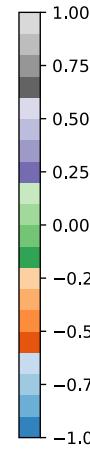
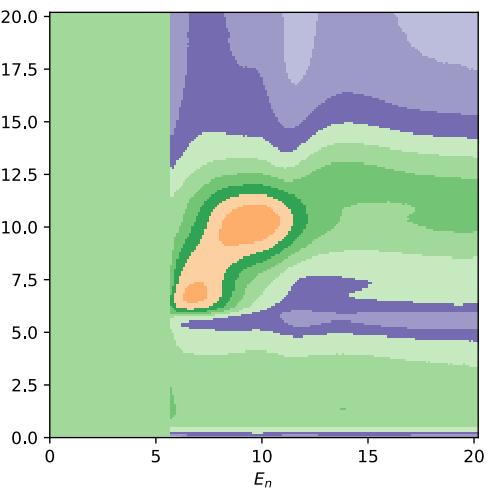
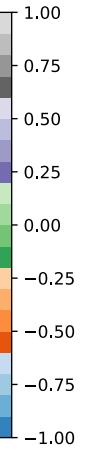
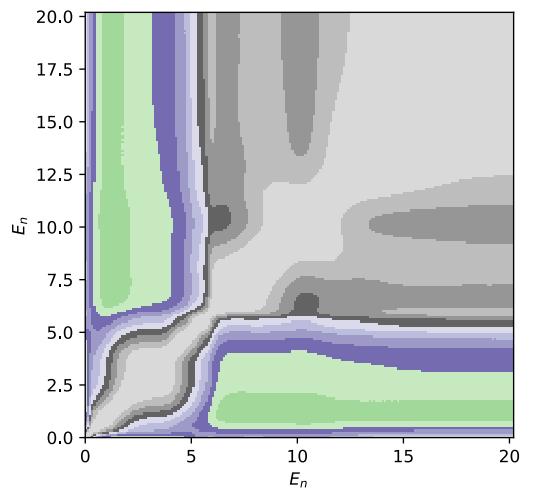
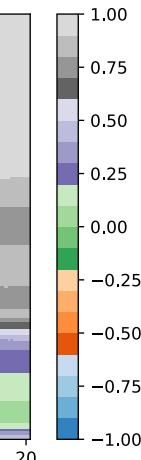
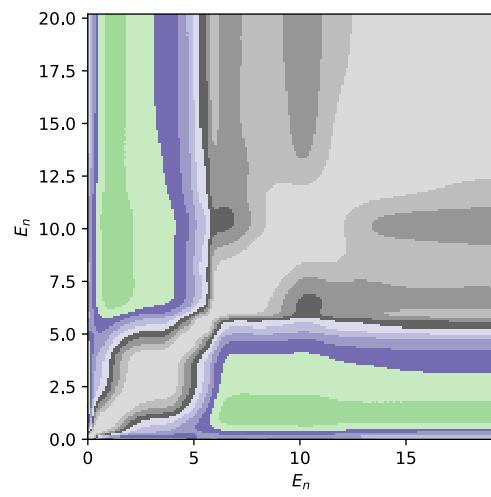
Other channels looking decent, uncertainty seems to be driven by fission barrier parameters.



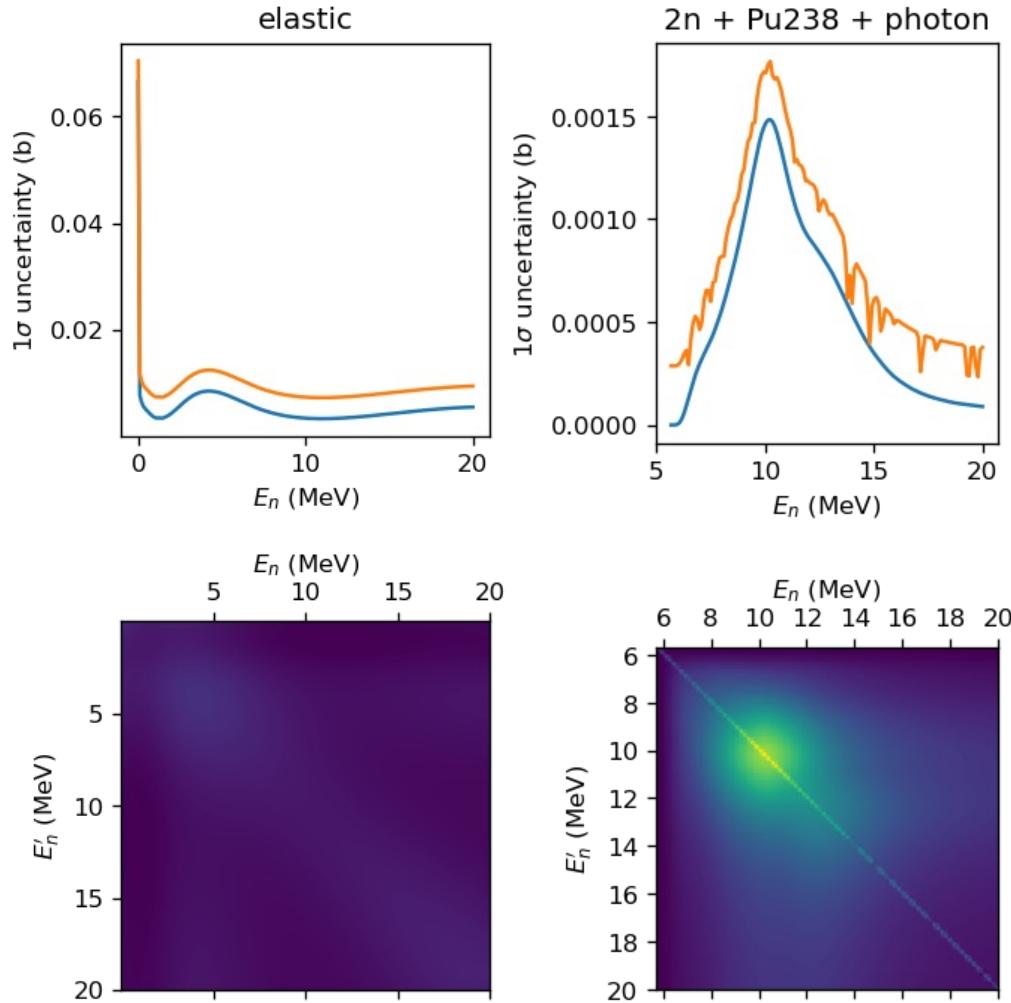
(n,f) vs (n,2n)



(n,el) vs (n,f)



Initial use of Gaussian Processes to soften model stiffness



- Model based (i.e. HF based) covariances can be stiff (small) due to stiffness and a lack of expressivity in the model.
- We want to extend the HF model to a super model
$$\sigma_\alpha(E) = \sigma_\alpha^{\text{HF}}(E) + \delta\sigma_\alpha^{\text{GPR}}(E)$$
- $\sigma_\alpha^{\text{HF}}(E)$ is defined by mean and covariance of the HF model
$$\sigma_\alpha^{\text{HF}}(E) \sim N(\mu_\alpha^{\text{HF}}, \Sigma_{\alpha\beta}^{\text{HF}})$$
- $\delta\sigma_\alpha^{\text{GPR}}(E)$ is from a Gaussian process, with mean zero and a data + model informed covariance
$$\delta\sigma_\alpha^{\text{GPR}}(E) \sim GPR(0, K_{\alpha\beta}(E, E'))$$
- We use a multi channel form for the kernel

$$K_{\alpha\beta}(E, E') = (w_\alpha w_\beta + \delta_{\alpha,\beta} \kappa_\alpha^2) e^{-\frac{(E-E')^2}{2l_\alpha l_\beta}}$$

Conclusions & future work

- Quick evaluation of covariances relies on parallel Monte Carlo approach.
- Incorporating new experimental data is relatively painless in the BFMC approach.
- GPs allow for the evaluation of improved means and covariances by absorbing some of the model uncertainty.
- Estimation of sensitivity will lead to reduced number of dimensions for more efficient MCMC sampling.
- Improved covariances from MCMC will be topped with the same GP approach as described here.
- No optical model (σ_{tot}) UQ yet...