

Peelle's Pertinent Puzzle and D'Agostini Bias — *Estimating the Mean with Relative Systematic Uncertainty*

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Peelle's Pertinent Puzzle (PPP)

• In 1987 at Oak Ridge National Lab, R.W. Peelle described having two observations to estimate a shared mean:

$$
y_1 = 1.0 \pm 10\% \qquad y_2 = 1.5 \pm 10\%
$$

and 20% fully correlated uncertainty

- Uncertainties
	- − "Fully correlated" means multiplicatively up or down together.
	- − Three uncertainties are taken to be independent.
	- − Relative uncertainties are assumed to be "1 sigma" values
- Peelle was a physicist doing nuclear data evaluation and he worked out the generalized least squares (GLS) estimate of the mean, as is standard in the field

Peelle's Pertinent Puzzle (PPP)

• Generalized least squares as per Peelle:

$$
y = (1.0, 1.5)^T
$$

\n $\hat{\mu} = (\mathbf{1}^T \Sigma^{-1} \mathbf{1})^{-1} \mathbf{1}^T \Sigma^{-1} y$
\n $\hat{\sigma}^2 = (\mathbf{1}^T \Sigma^{-1} \mathbf{1})^{-1}$
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• And the answer is ...

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$$

\n
$$
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$$

• And the answer is \ldots

 $\widehat{\mu} = 0.88 \pm 0.25$

below **both** observations!

- Terminology: *Data Weighted – Generalized Least Squares (DW-GLS)* refers to GLS with y in the assumed covariance matrix
	- − Many scientists don't consider having y in its own covariance as odd

As fully-correlated error increases we expect

- *proportional* **increase in uncertainty,**
- *and not* **additional bias…**

… but neither holds with DW-GLS

This problem exists more widely when using generalized least squares for nuclear data evaluation

Fitted curves (prior not shown)

- **Nominal case** has 0.42% fully-correlated error and looks OK but
	- − fitted is 0.7% below data
	- − 73% of residuals are positive
- Increasing the fully-correlated error by **4x** and **16x** reveals dramatic bias

High-Energy Physicists know this as *D'Agostini Bias*

Fig. 2. R measurements from PETRA and PEP experiments with the best fits of $OED + OCD$ to all the data (full line) and only below 36 GeV (dashed line). All data points are correlated (see text).

Example of a recent non-explanation

> *"The best-fit undershoots the data, essentially because with multiplicative uncertainties a lower prediction has a smaller uncertainty [D'Agostini, 1994]."*

D'Agostini, Giulio. "On the use of the covariance matrix to fit correlated data." *Nuclear Instruments and Methods in Physics Research Section A: Accelerators, Spectrometers, Detectors and Associated Equipment* 346.1-2 (**1994**): 306-311.

Ball, Richard D., et al. "Precision determination of the strong coupling constant within a global PDF analysis: NNPDF Collaboration." *The European Physical Journal* C 78 (**2018**): 1-16.

What is going on in PPP?

- **Is this not a problem at all?**
	- − Two observations with correlated errors aren't at all unlikely to fall on the same side of the mean
		- Why biased low rather than high? Symmetry would seem to make a case for no preference in either direction
		- **Why so far outside the data? The estimate makes one observation a** 3σ **outlier**
- Is this a problem with the GLS estimator for relative, correlated errors?
- Is this a fluke of the two-observation problem?

What is going on here?

- Is this not a problem at all?
	- − Two observations with correlated errors aren't at all unlikely to fall on the same side of the mean
- **Is this a problem with the GLS estimator for relative, correlated errors?**
	- − Let's explore a couple other estimators:
		- Maximum likelihood estimator: $y \sim \text{MVN}(\mu, \text{diag}(0.1^2 \mu^2) + 0.2^2 \mu \mu^T)$
		- Iteratively re-weighted least squares:

Do GLS, but iterate from initial $\hat{\mu} = y$ using $\Sigma = \text{diag}(0.1^2 \ \hat{\boldsymbol{\mu}}^2) + 0.2^2 \ \hat{\boldsymbol{\mu}} \hat{\boldsymbol{\mu}}^T$

• Is this a fluke of the two-observation problem?

What is going on here?

- Is this not a problem at all?
	- − Two observations with correlated errors aren't at all unlikely to fall on the same side of the mean

 $\hat{\mu}_{IRIS} = 1.25$

 $\hat{\mu}_{MIF} = 1.531$

- **Is this a problem with the GLS estimator for relative, correlated errors?**
	- − Let's explore a couple other estimators:
		- Maximum likelihood estimator:
		- Iteratively re-weighted least squares:
- Is this a fluke of the two-observation problem?

Recall $y_1 = 1.0 \pm 10\%$ $y_2 = 1.5 \pm 10\%$

MLE is **above** both observations!

Observations just touch at $+2\sigma$ MLE links the mean to this spread.

What is going on here?

- Is this not a problem at all?
	- − Two observations with correlated errors aren't at all unlikely to fall on the same side of the mean
- Is this a problem with the GLS estimator for relative, correlated errors?
- **Is this a fluke of the two-observation problem?**
	- − Simulate 500 samples, each with
		- *n*=100 observations
		- \blacksquare mean=1.25
		- SDs: 10% independent + 20% fully-correlated

Estimates of the mean parameter from 500 simulated samples

Now the MLE looks great!

> Why does the MLE look great now?

Estimates of the mean parameter with data generated with 12% independent error when 10% is assumed

The MLE no longer looks so great

> The MLE is sensitive to mis-specified

Conclusions so far:

- DW-GLS (i.e., using σ^2 yy^T in the covariance) biases estimates toward zero
	- − large $\sigma \Rightarrow$ bad bias.
- The MLE (and Bayes) is efficient but sensitive to mis-specified uncertainties
- IRLS is unbiased and robust to mis-specified uncertainties

PPP / D'Agostini Bias features in many papers over 30 years

In fact, IRLS has been reinvented to fix the bias

− but it is not recognized as fully legit

E.g., from Capote, *et al.* (2009)

> *An empirical "fix" compensates for PPP in a practical way as suggested by Chiba and Smith (1991).*

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Why does DW-GLS behave so badly?

Insight from a regression problem

Regress y on two predictors

$$
y = \alpha \mathbf{1} + \beta x + \gamma z + \epsilon
$$

Suppose we care about α , β , but γ is just a nuisance with priors

$$
\gamma \sim N(0, \sigma_{\gamma}^2), \qquad \epsilon \sim N(0, \sigma_{\epsilon}^2 I)
$$

Integrating over γ :

 $y = \alpha 1 + \beta x + e$ $e \sim \text{N} (0, \sigma_{\epsilon}^2 I + \sigma_{\gamma}^2 ZZ^T)$

I.e., one-predictor but correlated errors

Application to DW-GLS

DW-GLS uses $\Sigma=0.1^2\cdot\text{diag}(\mathbf{y}^2)+0.2^2\cdot\mathbf{y}\mathbf{y}^{\text{T}}$

and the $yy^{\rm T}$ is like regressing y on itself!

With $z = y$, the covariance absorbs the signal and

$$
(\hat{\alpha}\hat{\beta}) \rightarrow (0,0) \text{ as } \sigma_{\lambda} \rightarrow \infty
$$

Larger correlated error \Rightarrow more over-fitting

That is, DW-GLS with a fully correlated term yy^T is inherently biased toward zero.

Coming back to DW-GLS, is it at least OK for small correlated uncertainty?

No

- − It is not consistent.
- − Adding more data drives the estimate to 0 for *any* relative correlated uncertainty, *p>0*

Conclusions

Advice for analysts

- DO NOT represent fully-correlated relative uncertainty as p^2yy^T
	- $-$ equivalent to adding y as a predictor for y itself
	- − this drives the signal of interest to zero
	- − Results in bad mean AND bad uncertainty
- DO represent the uncertainty as p^2 $\widehat{\mu}$ $\widehat{\mu}^T$ and use IRLS
	- $-$ this is robust to mis-specified p
	- − also robust to mis-specified independent error
- **DO** carefully consider how relative error can allow noise to influence the mean estimates with MLE and Bayesian inference

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