

#### Peelle's Pertinent Puzzle and D'Agostini Bias — Estimating the Mean with Relative Systematic Uncertainty

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#### **Peelle's Pertinent Puzzle (PPP)**

• In 1987 at Oak Ridge National Lab, R.W. Peelle described having two observations to estimate a shared mean:

$$y_1 = 1.0 \pm 10\%$$
  $y_2 = 1.5 \pm 10\%$ 

and 20% fully correlated uncertainty

- Uncertainties
  - "Fully correlated" means multiplicatively up or down together.
  - Three uncertainties are taken to be independent.
  - Relative uncertainties are assumed to be "1 sigma" values
- Peelle was a physicist doing nuclear data evaluation and he worked out the generalized least squares (GLS) estimate of the mean, as is standard in the field



#### **Peelle's Pertinent Puzzle (PPP)**

• Generalized least squares as per Peelle:

$$y = (1.0, 1.5)^{T} \qquad \qquad \hat{\mu} = (\mathbf{1}^{T} \Sigma^{-1} \mathbf{1})^{-1} \mathbf{1}^{T} \Sigma^{-1} y$$
  
$$\Sigma = 0.1^{2} \cdot \text{diag}(y^{2}) + 0.2^{2} \cdot y y^{T} \qquad \qquad \hat{\sigma}^{2} = (\mathbf{1}^{T} \Sigma^{-1} \mathbf{1})^{-1}$$

• And the answer is ...



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$$\boldsymbol{\Sigma} = 0.1^2 \cdot \operatorname{diag}(\boldsymbol{y}^2) + 0.2^2 \cdot \boldsymbol{y} \boldsymbol{y}^T \qquad \qquad \hat{\sigma}^2 = (\mathbf{1}^T \boldsymbol{\Sigma}^{-1} \mathbf{1})^{-1}$$

And the answer is ....

 $\hat{\mu} = 0.88 \pm 0.25$ 

below **both** observations!

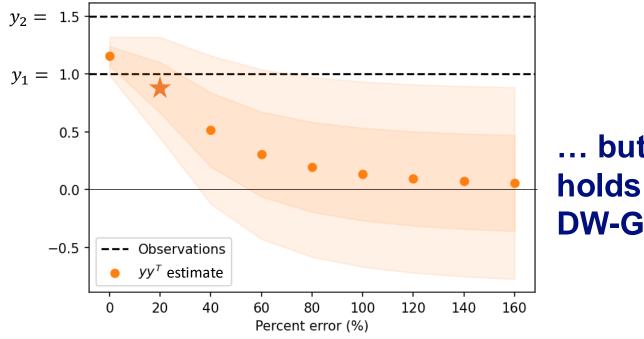
- Terminology: Data Weighted Generalized Least Squares (DW-GLS) refers to GLS with y in the assumed covariance matrix
  - Many scientists don't consider having y in its own covariance as odd



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#### As fully-correlated error increases we expect

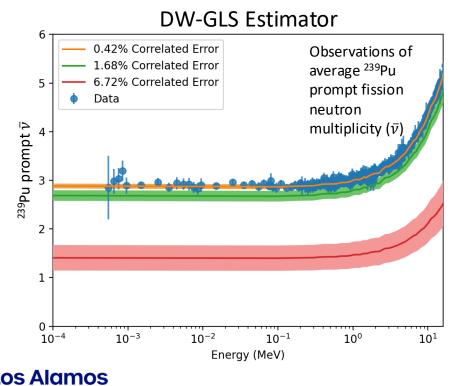
- proportional increase in uncertainty,
- and not additional bias...



... but neither holds with DW-GLS



# This problem exists more widely when using generalized least squares for nuclear data evaluation



Fitted curves (prior not shown)

- Nominal case has 0.42% fully-correlated error and looks OK but
  - fitted is 0.7% below data
  - 73% of residuals are positive
- Increasing the fully-correlated error by 4x and 16x reveals dramatic bias

### High-Energy Physicists know this as *D'Agostini Bias*

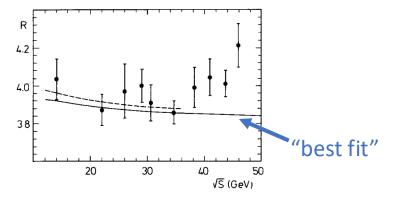


Fig. 2. R measurements from PETRA and PEP experiments with the best fits of QED+QCD to all the data (full line) and only below 36 GeV (dashed line). All data points are correlated (see text). Example of a recent non-explanation

"The best-fit undershoots the data, essentially because with multiplicative uncertainties a lower prediction has a smaller uncertainty [D'Agostini, 1994]."

**D'Agostini,** Giulio. "On the use of the covariance matrix to fit correlated data." *Nuclear Instruments and Methods in Physics Research Section A: Accelerators, Spectrometers, Detectors and Associated Equipment* 346.1-2 (**1994**): 306-311.

**Ball**, Richard D., et al. "Precision determination of the strong coupling constant within a global PDF analysis: NNPDF Collaboration." *The European Physical Journal* C 78 (**2018**): 1-16.

#### What is going on in PPP?

- Is this not a problem at all?
  - Two observations with correlated errors aren't at all unlikely to fall on the same side of the mean
    - Why biased low rather than high? Symmetry would seem to make a case for no preference in either direction
    - Why so far outside the data? The estimate makes one observation a  $3\sigma$  outlier
- Is this a problem with the GLS estimator for relative, correlated errors?
- Is this a fluke of the two-observation problem?



#### What is going on here?

- Is this not a problem at all?
  - Two observations with correlated errors aren't at all unlikely to fall on the same side of the mean
- Is this a problem with the GLS estimator for relative, correlated errors?
  - Let's explore a couple other estimators:
    - Maximum likelihood estimator:  $y \sim MVN(\mu, diag(0.1^2 \mu^2) + 0.2^2 \mu \mu^T)$
    - Iteratively re-weighted least squares:

Do GLS, but iterate from initial  $\hat{\mu} = y$  using  $\Sigma = \text{diag}(0.1^2 \ \hat{\mu}^2) + 0.2^2 \ \hat{\mu}\hat{\mu}^T$ 

• Is this a fluke of the two-observation problem?



#### What is going on here?

- Is this not a problem at all?
  - Two observations with correlated errors aren't at all unlikely to fall on the same side of the mean

 $\hat{\mu}_{MLE} = 1.531$ 

 $\hat{\mu}_{IRIS} = 1.25$ 

- Is this a problem with the GLS estimator for relative, correlated errors?
  - Let's explore a couple other estimators:
    - Maximum likelihood estimator:
    - Iteratively re-weighted least squares:
- Is this a fluke of the two-observation problem?

Observations just touch at  $\pm 2\sigma$ MLE links the mean to this spread.

MLE is **above** both

observations!

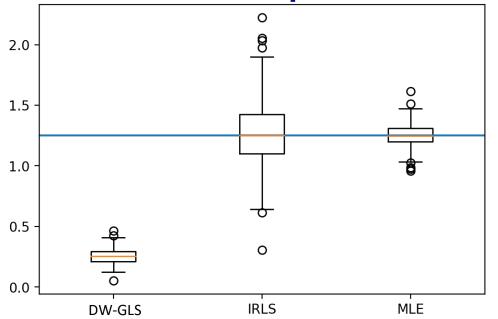
Recall  $y_1 = 1.0 \pm 10\%$  $y_2 = 1.5 \pm 10\%$ 

### What is going on here?

- Is this not a problem at all?
  - Two observations with correlated errors aren't at all unlikely to fall on the same side of the mean
- Is this a problem with the GLS estimator for relative, correlated errors?
- Is this a fluke of the two-observation problem?
  - Simulate 500 samples, each with
    - n=100 observations
    - mean=1.25
    - SDs: 10% independent + 20% fully-correlated



### Estimates of the mean parameter from 500 simulated samples

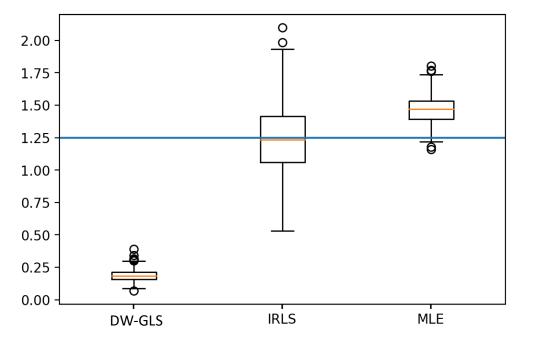


Now the MLE looks great!

Why does the MLE look great now?



### Estimates of the mean parameter with data generated with 12% independent error when 10% is assumed



The MLE no longer looks so great

> The MLE is sensitive to mis-specified uncertainty!



#### **Conclusions so far:**

- DW-GLS (i.e., using  $\sigma^2 y y^T$  in the covariance) biases estimates toward zero
  - large  $\sigma \Longrightarrow$  bad bias.
- The MLE (and Bayes) is efficient but sensitive to mis-specified uncertainties
- IRLS is unbiased and robust to mis-specified uncertainties



#### PPP / D'Agostini Bias features in many papers over 30 years

### In fact, IRLS has been reinvented to fix the bias

 but it is not recognized as fully legit

### E.g., from Capote, *et al.* (2009)

An empirical "fix" compensates for PPP in a practical way as suggested by Chiba and Smith (1991).



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#### Why does DW-GLS behave so badly?



# Insight from a regression problem

Regress y on two predictors

$$\mathbf{y} = \alpha \mathbf{1} + \beta \mathbf{x} + \gamma \mathbf{z} + \boldsymbol{\epsilon}$$

Suppose we care about  $\alpha$ ,  $\beta$ , but  $\gamma$  is just a nuisance with priors

$$\gamma \sim \mathrm{N}(0, \sigma_{\gamma}^2), \qquad \boldsymbol{\epsilon} \sim \mathrm{N}(\boldsymbol{0}, \sigma_{\boldsymbol{\epsilon}}^2 \boldsymbol{I})$$

Integrating over  $\gamma$ :

 $\mathbf{y} = \alpha \mathbf{1} + \beta \mathbf{x} + \mathbf{e} \\ \mathbf{e} \sim \mathrm{N} (\mathbf{0}, \ \sigma_{\epsilon}^{2} \mathbf{I} + \sigma_{\gamma}^{2} \mathbf{z} \mathbf{z}^{T})$ 

I.e., one-predictor but correlated errors

#### **Application to DW-GLS**

DW-GLS uses  $\Sigma = 0.1^2 \cdot \text{diag}(y^2) + 0.2^2 \cdot yy^T$ 

and the  $yy^{T}$  is like regressing y on itself!

With z = y, the covariance absorbs the signal and

$$(\hat{\alpha} \ \hat{\beta}) \rightarrow (0,0)$$
 as  $\sigma_{\lambda} \rightarrow \infty$ 

Larger correlated error  $\Rightarrow$  more over-fitting

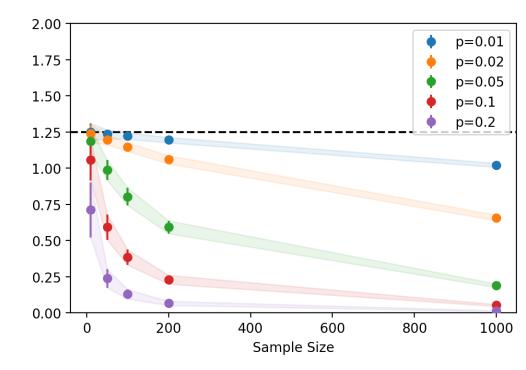
That is, DW-GLS with a fully correlated term  $yy^{T}$  is inherently biased toward zero.



# Coming back to DW-GLS, is it at least OK for small correlated uncertainty?

#### No

- It is not consistent.
- Adding more data drives the estimate to 0 for *any* relative correlated uncertainty, *p>0*





#### Conclusions

#### Advice for analysts

- **DO NOT** represent fully-correlated relative uncertainty as  $p^2yy^T$ 
  - equivalent to adding y as a predictor for y itself
  - this drives the signal of interest to zero
  - Results in bad mean AND bad uncertainty
- **DO** represent the uncertainty as  $p^2 \hat{\mu} \hat{\mu}^T$  and use IRLS
  - this is robust to mis-specified p
  - also robust to mis-specified independent error
- **DO** carefully consider how relative error can allow noise to influence the mean estimates with MLE and Bayesian inference



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